

# Uncertainty in Ecological Analysis

Noel Cressie

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and Environmental Sciences

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([www.stat.ohio-state.edu/~sses](http://www.stat.ohio-state.edu/~sses))



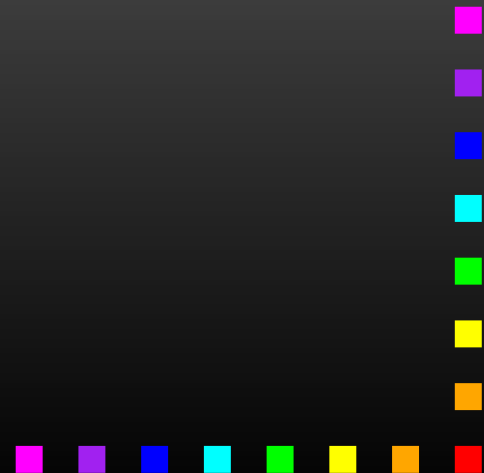
# The Workshop

- Science
  - MBI Organizing Committee, 2005-6 Year of Ecology and Evolution
  - Workshop Organizing Committee: Kate Calder, Jim Clark, Noel Cressie (Chair), Jay Ver Hoef, Chris Wikle
  - You (Ohio State, USA Academic/Gvt, Canada, UK, France, Iceland, Brazil, Australia, New Zealand)



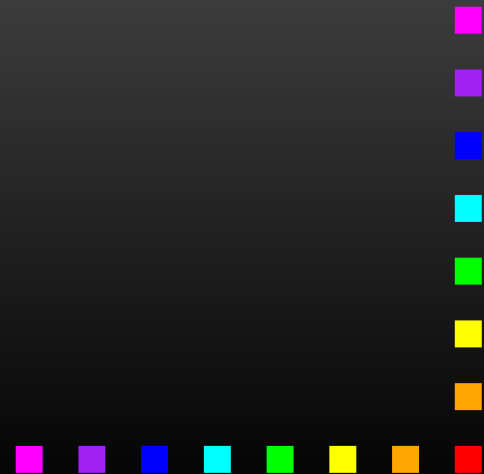
# The Workshop, ctd.

- Finances/Visas/Facilities/Computing (MBI)
  - Avner Friedman, Director
  - Tony Nance, Assistant Director
  - Nikki, Matt, Michael, Becca, Stella



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  - Tony Nance, Assistant Director
  - Nikki, Matt, Michael, Becca, Stella
- Everything Else (SSES)
  - Terry, Kate, Noel



# The Workshop, ctd.

- Discussion, Discussion, Discussion!
  - Shorter Talks (Speakers'/Chairs' responsibility)
  - 2 Discussants Per Session (3-5 mins. each)
  - Lots of Time for General Discussion
  - Ecological Study Discussion (6 separate groups)
  - Panel Session on Last Day
  - Healthy Debate Please!



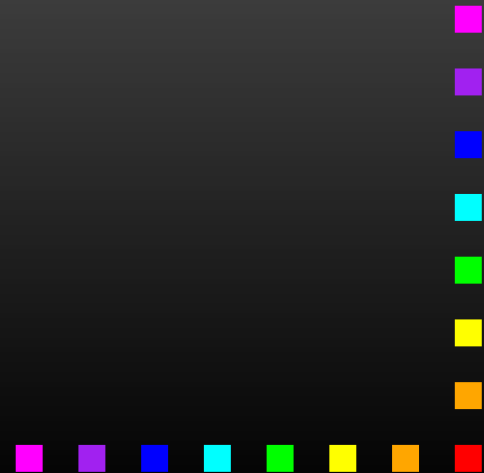
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- Party, Party, Party!
  - Opening Reception This Evening (5 p.m., 7th floor of Math Tower)
  - SSES Sponsored Reception on Tuesday (6-8:30 p.m., OSU Wetland)
  - Workshop Dinner on Wednesday (Holiday Inn: 6-7 p.m., Cash Bar; 7 p.m., Dinner)
  - Lunch/Reception on Thursday (Holiday Inn: 12 noon)



# Sources of Uncertainty

- Data, Data, Data!



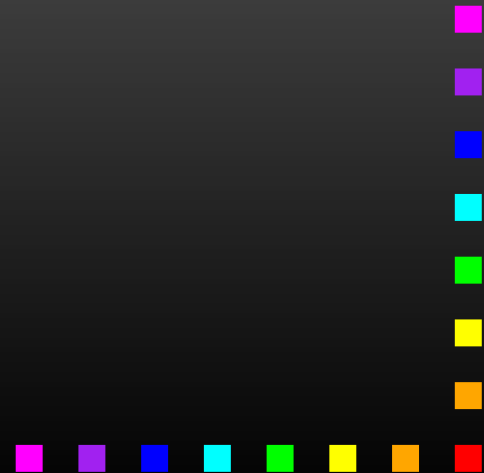
# Sources of Uncertainty

- Data, Data, Data!
- Models
  - Stochastic or deterministic
  - Stochastic: Some components modeled as random; allows simple description of incomplete science knowledge
  - Parametric classes of models

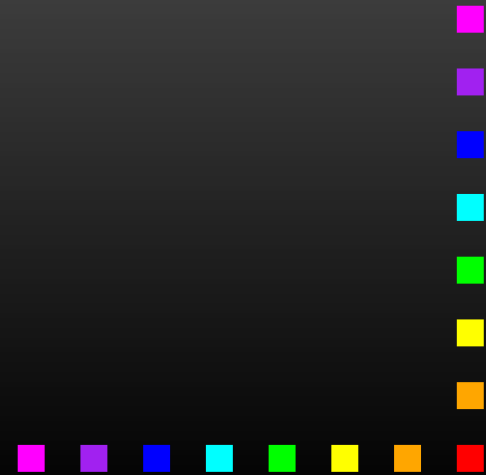


# Sources of Uncertainty, ctd.

- Parameters
  - Fixed (but unknown)?
  - Random?
  - We all want to make inference on unknown parameters, but how?

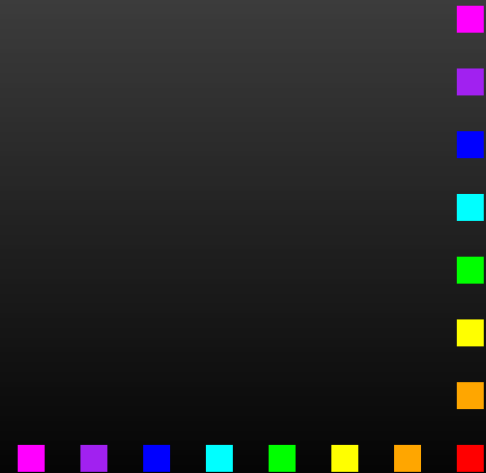


# Topics to Avoid at Parties (But not at this Workshop!)



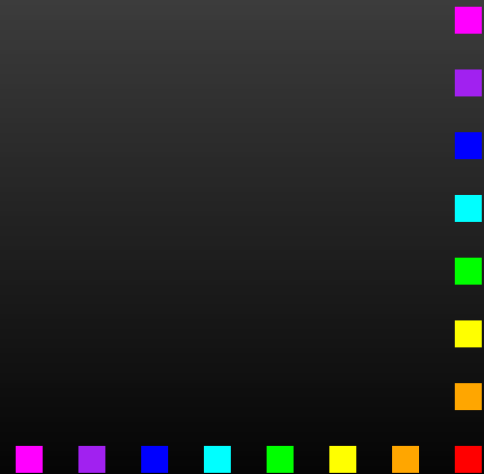
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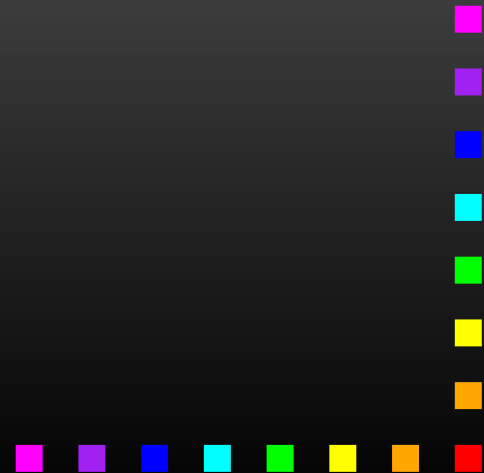
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- Model Parameters - Fixed or random?



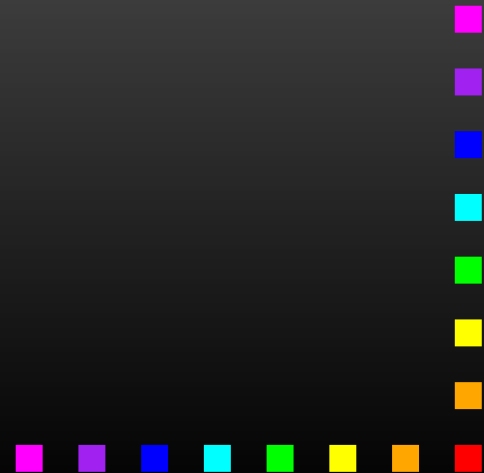
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  - Classical (Fixed Parameters): Byron Morgan
  - Bayesian (Random Parameters): Mark Berliner



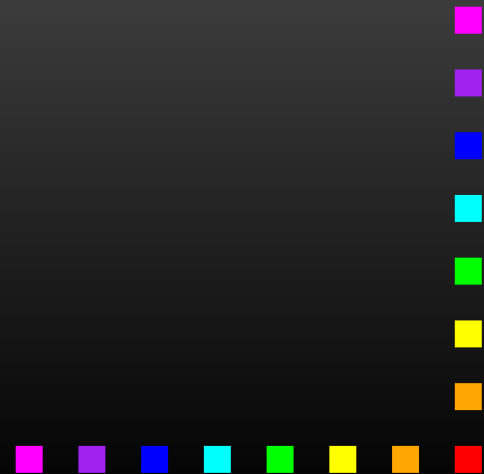
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- Philosophical - comes from experience with different types of problems and from beliefs!
- Tutorial-Style presentations this AM
  - Classical (Fixed Parameters): Byron Morgan
  - Bayesian (Random Parameters): Mark Berliner
- Group sessions on Mon, Tue, Wed
  - Ecological Study: Jay Ver Hoef
  - 6 Groups
  - Group Leaders summarize results (5 min. ea.) at Workshop Dinner on Wed.



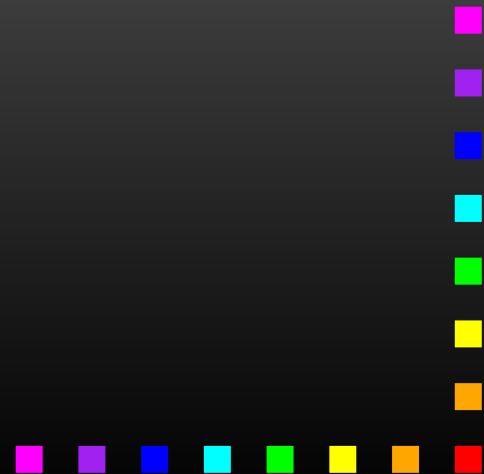
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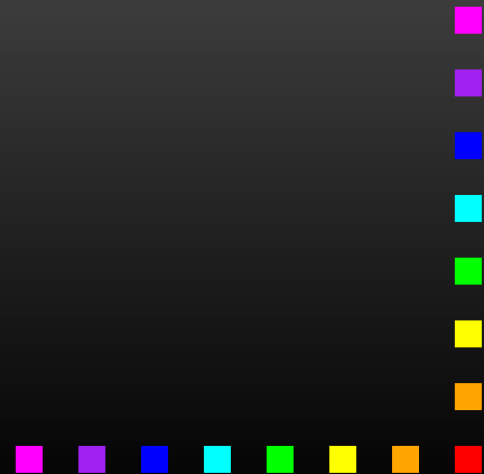
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- Earth could be viewed as the organism and we only have one replicate (to get it right)!



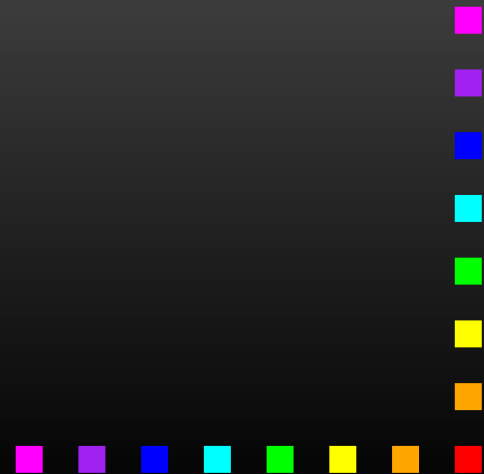
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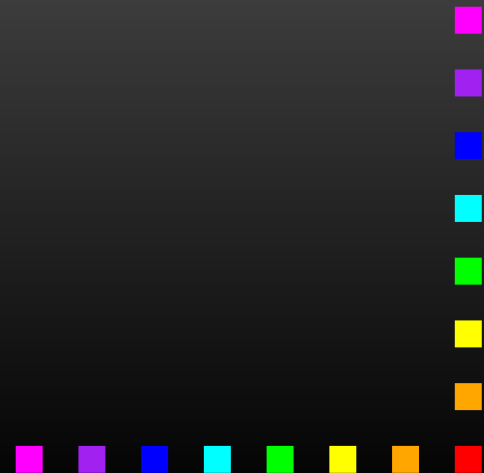
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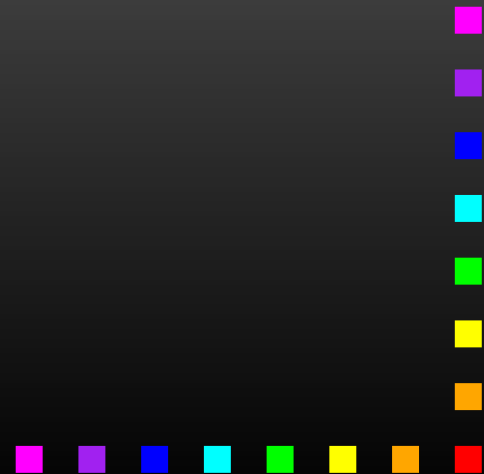
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- Sometimes one or other is aggregated
- Temporal - Dynamic models, expressing current states in terms of past states, allow explanation (causation) and forecasting
- Spatial - Dynamics can be spatially inhomogeneous! In other words, “why” can depend on “where”. (MBI Spatial Ecology Workshop was last month)



# Spatial Support

Uncertainty due to lack of knowledge. Some sources:

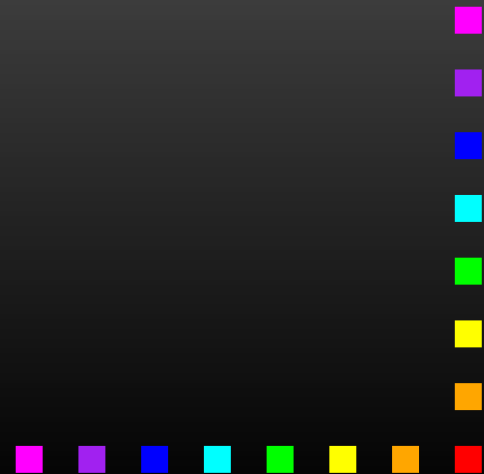
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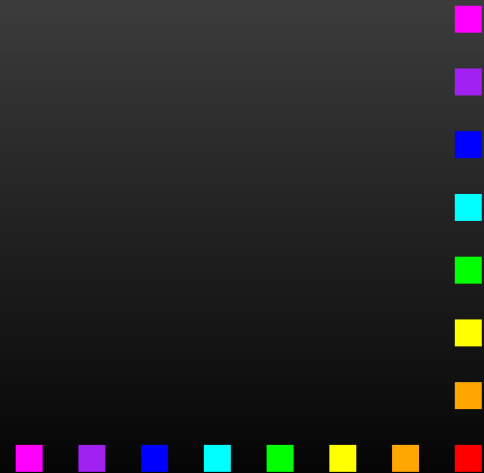
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- Incomplete data (missing variables and missing observations on variables)



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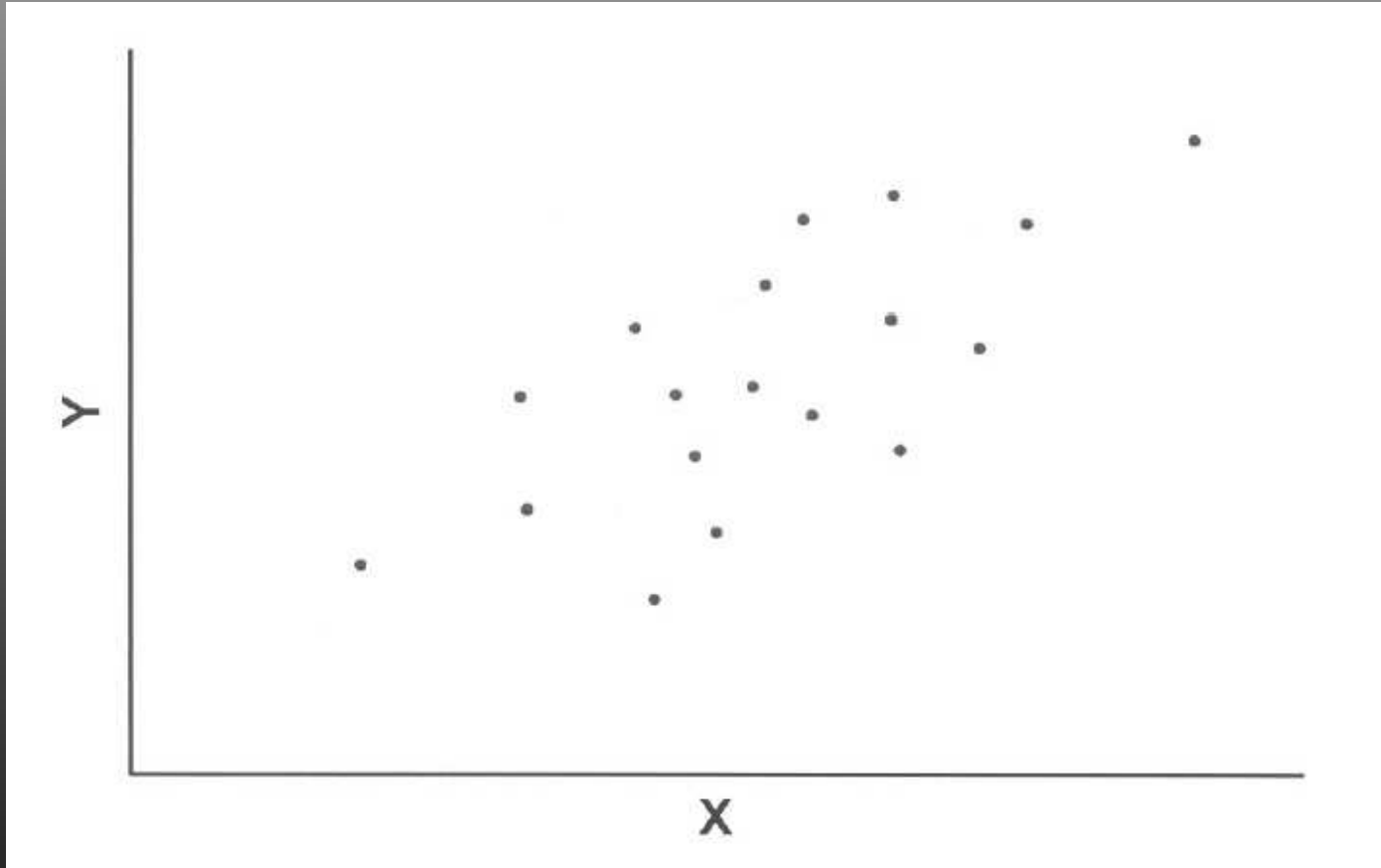
# Spatial Support

Uncertainty due to lack of knowledge. Some sources:

- Science (incomplete or incorrect models)
- Incomplete data (missing variables and missing observations on variables)
- Measurement (noise)
- Spatial uncertainty

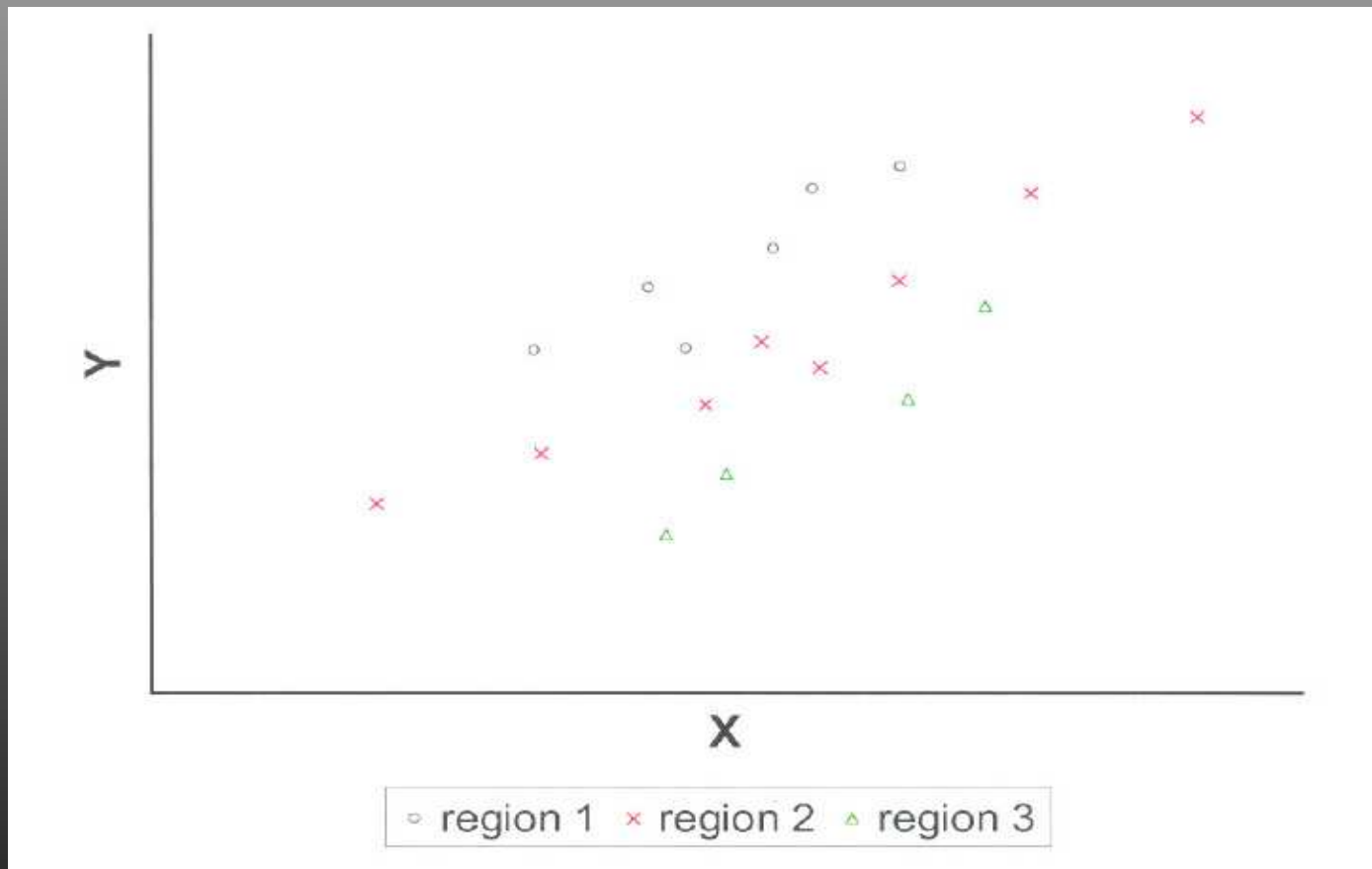
# Spatial Support, ctd.

- Spatial uncertainty
  - Object view versus field view
  - Location error
  - Aggregation (COS, MAUP, ecological bias)  
This is, in effect, Simpson's Paradox in a spatial setting



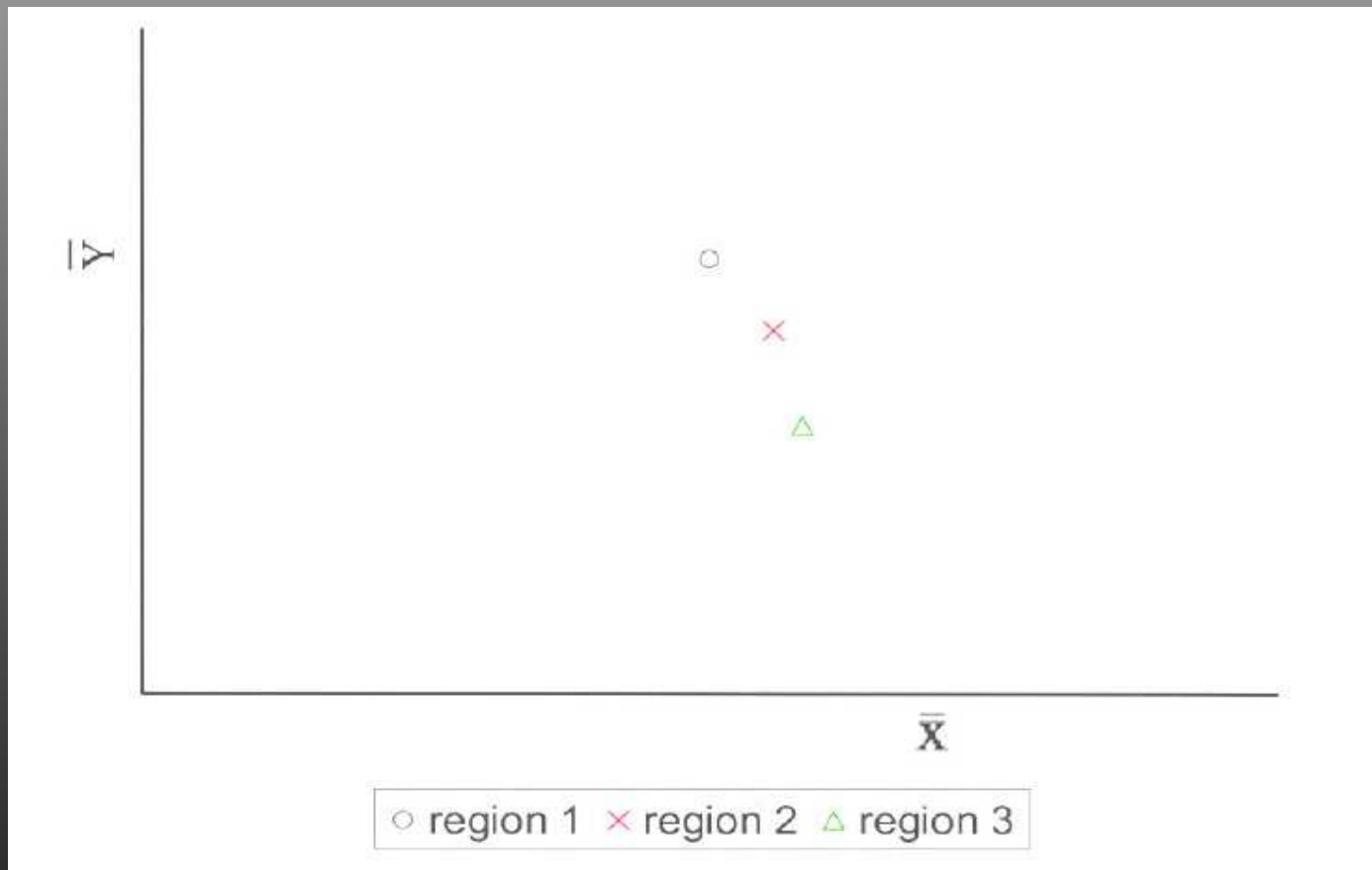
$$Y = \beta_0^* + \beta_1^* X + \varepsilon; \quad \hat{\beta}_1^* > 0$$





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$$\bar{Y} = \beta_0^* + \beta_1^* \bar{X} + \bar{\varepsilon}; \quad \hat{\beta}_1^* < 0$$



# Ecological Inference

Deducing *individual* behavior from *aggregate* data  
Robinson (1950) *American Sociological Review*:  
1930 data from lower 48 states:

$Z$  = % foreign born in a state

$Y$  = % literate in a state

$\Rightarrow$  corr. = 0.53

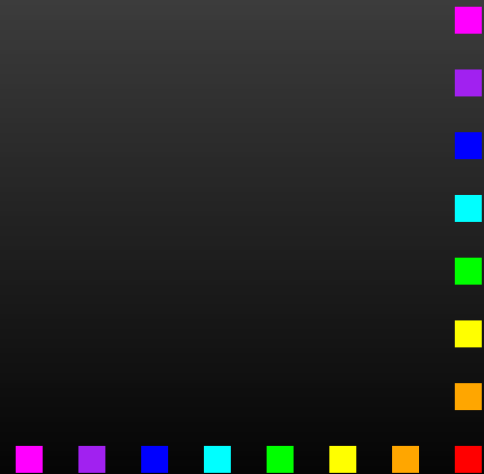
Based on sampling, Robinson found individual  
corr. =  $-0.11$

Thus, corr. value of 0.53 represents an *ecological bias*



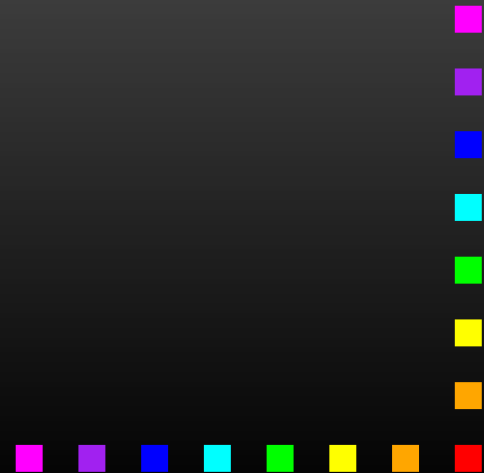
# Approach

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- Use statistical modeling to avoid ecological fallacy
- Fit to data that are aggregated (counts) using model built at point level
- Have to have a good point-level model!

# (Aggregated) Disease Counts

$Y(A_1), \dots, Y(A_n)$  modeled hierarchically:

$$Y(A_i) \mid p_i \sim \text{Poi}(N_i p_i), \quad \text{indep.}$$

$$\log p_i \mid x(A_i) = \log(E_i/N_i) + \beta_0 + \beta_1 x(A_i) + \delta_i,$$

where  $x(A_i)$  is an exposure variable

$\delta_i$  is a (spatial) error term.

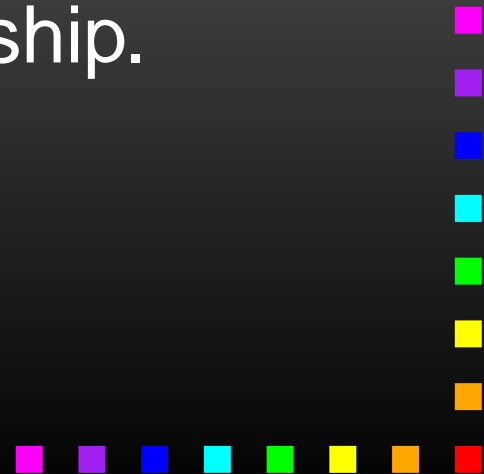
Finally,  $\{x(A_i)\}$  is potentially a spatial process.



# (Aggregated) Disease Counts, ctd.

Inference on  $\beta = (\beta_0, \beta_1)$  might not reflect *disease/exposure relationship*, which occurs at the individual level. We see the possibility of the *ecological fallacy*.

Let  $\beta^*$  denote *individual-level coefficients* involved in disease/exposure relationship.



# (Individual-Level) Disease Mechanism

$$Y_{ij} = \begin{cases} 1 & j\text{-th individual in area } A_i \text{ has disease} \\ 0 & \text{otherwise} \end{cases}$$

$$\begin{aligned} \Pr(Y_{ij} = 1) &= \int \Pr(Y_{ij} = 1 | x_{ij}, \beta^*) f(x_{ij} | \phi_i) dx_{ij} \\ &\equiv p(\beta^*, \phi_i), \end{aligned}$$

where  $f(\cdot | \phi_i)$  is the *within-area distribution* of exposure within  $A_i$ .



# Disease Mechanism (cont.)

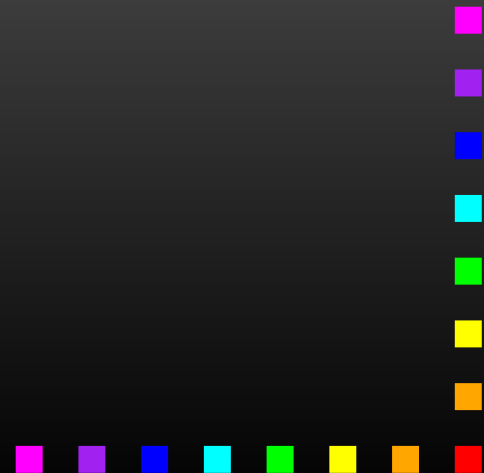
$$\Rightarrow Y_{ij} | \beta^*, \phi_i \sim \text{Bernoulli}(p(\beta^*, \phi_i))$$

$$\Rightarrow Y(A_i) | \beta^*, \phi_i \sim \text{Bin}(N_i, p(\beta^*, \phi_i))$$

$$\Rightarrow Y(A_i) \dot{\sim} \text{Poi}(N_i p(\beta^*, \phi_i))$$

That is,

$$Y(A_i) \dot{\sim} \text{Poi}(\lambda(A_i; \psi))$$



# Loglinear Individual-Level Model

$$\begin{aligned}\Pr(Y_{ij} = 1 | x_{ij}, \beta^*) &= \exp(\beta_{0,i}^* + \beta_1^* x_{ij}) \\ \Rightarrow p(\beta^*, \phi_i) &= \exp(\beta_{0,i}^*) \int \exp(\beta_1^* x_{ij}) f(x_{ij} | \phi_i) dx_{ij} \\ &\neq \exp(\beta_{0,i}^*) \exp(\beta_1^* \int x_{ij} f(x_{ij} | \phi_i) dx_{ij}) \\ &\equiv \exp(\beta_{0,i}^* + \beta_1^* x(A_i)),\end{aligned}$$

Clearly, the use of sample-survey estimated covariate  $\{\hat{x}(A_i)\}$  results in an *ecological bias*.



# Likelihood based on case counts

- Data are counts

$$(Y(A_1), \dots, Y(A_n))$$

- Same intensity and Poisson assumptions
- Likelihood of  $\psi$ :

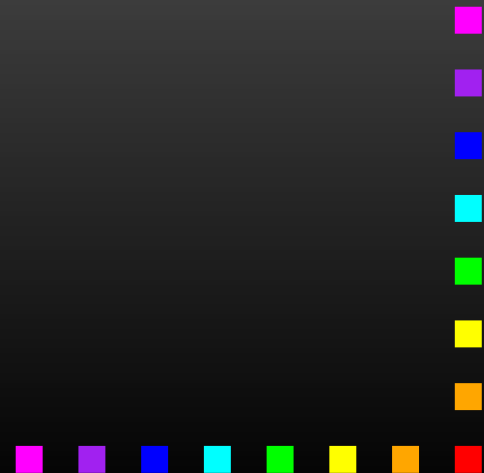
$$\ell(\psi) = \prod_{i=1}^n \frac{e^{-\lambda(A_i; \psi)} (\lambda(A_i; \psi))^{Y(A_i)}}{Y(A_i)!}$$



# Likelihood based on case counts, ctd.

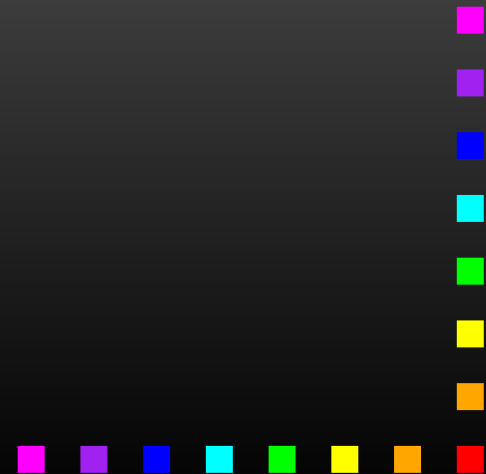
Parameters  $\psi = (\beta_1^*, \{\beta_{0,i}^*\}, \{\phi_i\})$  are defined at individual level but inference is based on aggregated data (counts),

$$(Y(A_1), \dots, Y(A_n))$$



# Ecological Analysis

This workshop is about



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This workshop is about  
... trying to be certain about the ecological  
problem's uncertainties ...

